Associations between peer attachment and neural correlates of risk processing across adolescence

J. Susanne Asscheman, Kirby Deater-Deckard, Nina Lauharatanahirun, Susanne Koot, Brooks King-Casas, Jungeun Kim-Spoon

1. Introduction

Adolescence is a period of increased risk-taking behavior where individual differences in risk taking may relate to both adverse and positive experiences with peers. Yet, knowledge on how risk processing develops in the adolescent brain and whether this development is related to peer attachment is limited. In this longitudinal functional magnetic resonance imaging (fMRI) study, we collected data from 167 adolescents (53% male) followed for four annual assessments across ages 13–17 years. At each assessment, participants completed a lottery choice task to assess neural risk processing and reported on their perceived attachment to peers and parents. Behaviorally, risk-preference on the lottery choice task decreased linearly with age. Neural activation during risk processing was consistently found in the insula and dACC across the four assessments and increased linearly from ages 13–17 years. Furthermore, higher peer attachment was related to greater right insula risk processing for males but not for females, even after controlling for parental attachment. The magnitudes of this association did not change with age. Findings demonstrate that neural risk processing shows maturation across adolescence and high peer attachment may be associated with low risk taking by heightening neural sensitivity to potential risks for male adolescents.
et al., 2015; van Duijvenvoorde et al., 2016), but another important process that may guide risk-taking behavior is the neural evaluation of risks (d’Acremont and Bossaerts, 2008; van Duijvenvoorde et al., 2015).

However, research is still limited on the neural development of the latter process during adolescence.

Two regions that may be especially important for risk processing and guiding adolescent risk behavior are the insula and dACC. Extant literature in adults and adolescents implicates the insular cortex and dACC in risk processing (see meta-analysis by Mohr et al., 2016; Platt and Huettel, 2008; Schongen et al., 2011). Longitudinal studies on the development of risk processing in the insula and dACC are lacking. Some cross-sectional neuroimaging studies found that, during risky decisions (e.g., choosing risky options over safe options), adults showed higher insula and dACC activation compared to adolescents and children (Paulsen et al., 2012). In contrast, other cross-sectional neuroimaging studies found an age-related decrease in medial PFC/dACC activation for high-risk choices (compared to low-risk choices) among children, adolescents, and adults (van Leijenhorst et al., 2006, 2010). Finally, one study examining neural sensitivity to risk information (i.e., outcome variability) found that insula and dmPFC/d ACC activation were hyperactive in adolescents compared to children and adults, and that higher activation in the insula and dmPFC/dACC was related to greater risk-averse behavior among adolescents (van Duijvenvoorde et al., 2015). These results suggest the importance of the insula and dACC for guiding risk-related decision making particularly among adolescents.

However, because of the cross-sectional nature, prior studies did not assess within-person changes in risk-related neural processing and were limited in making developmental inferences. The mixed findings reported in the current literature may be in part due to of the fact that prior cross-sectional studies compared different age groups: different studies had different age ranges to represent children, adolescent, and adult groups, and only small numbers of individuals were included in each age group (i.e., 12–25 people per age group). In the current study, we examined the development of neural risk processing by studying within-person changes in the insula and dACC across adolescence using a relatively large sample size of adolescents who were assessed repeatedly from 13 to 17 years.

Besides increased risk-taking behavior, adolescence is also characterized by an increased sensitivity to the peer environment (Nelson et al., 2005). As such, the peer environment may contribute to the development of risk-taking behavior and associated neural circuitries (Burnett et al., 2011; Schriber and Guyer, 2016). For example, prior studies showed that the presence of peers and winning a risky gamble for friends were linked to increased activation in brain regions related to reward processing (i.e., striatum activation) during risk-taking (Braams et al., 2014; Chein et al., 2011; Powers et al., 2018). Also, the quality of peer relationships, such as the level of attachment to peers, has been shown to be positively related to adolescent risk-taking behaviors (Piko, 2000; Urberg et al., 2003; Wills et al., 2004) and to be linked to increased activation in brain regions associated with the socio-emotional system (e.g., medial prefrontal cortex) during risk-taking (Telzer et al., 2015). Research also found an association between high friendship quality and increased activation in the striatum when winning a gamble for friends, but only for female adolescents (Braams et al., 2014) but this association between high friendship quality and increased activation in the striatum was not found consistently (Braams and Crone, 2016, 2017).

Thus, studies indicate the link between peer environment and heightened neural activation related to reward. It is unknown, however, whether positive peer experiences such peer attachment may be associated with adolescents’ neural sensitivity to risk. Prior studies suggest two opposing effects that peer attachment may have on neural risk processing. On the one hand, a prior study found that high levels of peer attachment may protect adolescents from risky behavior (Telzer et al., 2015). Accordingly, high levels of peer attachment may heighten adolescents’ sensitivity to risk information, ultimately resulting in low risk-taking behaviors. On the other hand, high levels of peer attachment may increase risk-seeking tendencies because peers may heighten the positive affect among adolescents. Adolescents high in peer attachment might experience higher positive affect from peers, compared to adolescents low in peer attachment, which may decrease the perceived riskiness of decisions (Romer and Hennessy, 2007). Furthermore, adolescents with high peer attachment may engage in greater peer socialization of risky behavior via social learning, compared to their counterparts with low peer attachment (Brechwald and Prinstein, 2011; Jessor, 1993; Kendell, 1985). As a result, higher peer attachment may be associated with lower neural sensitivity to risk information. It is difficult to ascertain which effect is more likely, because the nature of the tasks used in prior studies examining peer influences on adolescent brain functioning were not able to separate risk processing from reward processing. Moreover, prior studies were limited in their ability to evaluate developmental changes in peer influences on neural processing due to the cross-sectional nature of their fMRI data.

When studying the association between peer attachment and neural risk processing during adolescence, potential sex differences also need to be considered. To our knowledge, no prior study has examined sex differences in the associations between peer attachment and neural processing of risk information. However, one prior study found a positive association between friendship quality and neural reward processing during risk choices involving friends for females only (Braams et al., 2014). Moreover, sex differences in peer attachment experiences are known to exist. Females value relationship intimacy more, seek more support from peers, and disclose more information to their peers than males (Gorrese and Ruggieri, 2012; Rose and Rudolph, 2006). These sex differences may result in differences in how adolescents’ peer attachment is associated with the processing of risk information in their brains. Given the lack of prior empirical studies, we explored adolescent sex as a statistical moderator in the present study.

The current study had two main goals. First, we examined the development of neural risk processing and tested potential sex differences across adolescence, from ages 13–17 years. We focused on the insula and dACC based on extensive literature implicating these regions in risk processing. We hypothesized that risk processing activity in the insula and dACC would increase with age, based on previous cross-sectional studies showing both higher activation during risk processing in adults compared to adolescents and children (Eshel et al., 2007; Paulsen et al., 2012) and a suggested increase in risk-related activation in these regions among adolescents compared to children (van Duijvenvoorde et al., 2015). In addition, we were interested in knowing whether risk processing is specific to the insula and dACC, or if it involves other brain regions that have been shown to activate during risk-taking behavior such as the ventral striatum, orbitofrontal cortex, and lateral prefrontal cortex. Second, we further examined whether the development of neural risk processing was associated with peer attachment. Specifically, using longitudinal data spanning ages 13–17 years, we examined how peer attachment is associated with the development of neural processing of risk evaluation which was assessed using an economic risky decision-making task. In doing so, we additionally tested for a sex differences in the association between neural risk processing and peer attachment.

2. Method

2.1. Participants

The sample consisted of 167 adolescents (53% males) from rural-to-urban areas in an Appalachian region of southwestern Virginia in the United States of America. Data were collected on four annual assessments in 2014 (T1; Mage = 14.13 years, SD = 0.54), 2015 (T2; Mage = 15.05 years, SD = 0.54), 2016 (T3; Mage = 16.08 years, SD = 0.55), and 2017 (T4, Mage = 17.01 years, SD = 0.55). At T1, data from 157 adolescents were collected. Seventeen adolescents did not return at T2 due
to lost contact, declined participation or ineligibility for the task (i.e., brain abnormality, not meeting MRI safety criteria). Therefore, at T2 10 additional participants were included to reach a sample size of 167 adolescents. At T3 data from 147 participants, and at T4 data from 150 participants were collected. Participants who did not participate in one annual assessment were still invited to participate in following assessments. Overall, 24 adolescents did not participate at all possible time points for reasons including: ineligibility for tasks ($n = 2$), declined participation ($n = 17$), and lost contact ($n = 5$) during the follow-up assessments. Attrition analyses using general linear model (GLM) univariate procedures indicated that the rate of participation (indexed by proportion of years participated to years invited to participate) was not significantly predicted by demographic variables ($p = .61$ for age, $p = .67$ for income, $p = .62$ for sex, and $p = .73$ for race, contrasted as White vs non-White).

Adolescents were identified as 82% Caucasian, 12% African-American, and 6% other. The mean of the annual family household income fell between $25,000 and $34,999 at T1, which is below the median household income ($57,652) of the United States of America during the study period spanning the mid-2010s (Census, 2017). Ethnicity and household income of the current sample were representative for the area in which data were collected.

### 2.2. Procedure

Participants were recruited through advertisement using flyers, letters and emails. Adolescents with MRI contraindications (e.g., dental braces, history of head injuries) were excluded from participation. During each annual assessment, adolescents participated in a five-hour session at the university which was supervised by a trained research assistant. Participants filled out questionnaires and completed a two-hour MRI session in which an anatomical scan was acquired followed by several fMRI tasks of which the risk-taking task was the first. Parents were informed about the study and provided written informed consent while adolescents provided written assent. Adolescents were monetarily compensated at the end of each annual visit for their participation. All study procedures were approved by the Institutional Review Board of Virginia Tech.

### 2.3. Measures

#### 2.3.1. Peer attachment

Adolescents reported annually on their perceived attachment to peers and parents using a shortened version of the Inventory of Parent and Peer Attachment (IPPA; Armsden and Greenberg, 1987; Raja et al., 1992). The IPPA is designed to assess adolescents’ perceptions on the quality of psychological security provided by parents and peers as well as how accessible and responsive these attachment figures are. The shortened version consisted of 12 items on three subscales (4 items per scale): trust, communication, and alienation. The trust subscale assessed the degree of mutual trust parents and peers provide (e.g., “I can count on my friends when I need to get something off my chest.”). The communication subscale assessed the perceived quality of communication between adolescents and their parents or peers (e.g., “When we discuss things, my friends care about my point of view.”). The alienation subscale assessed the extent of anger or alienation adolescents experience in their relationship with parents or peers (e.g., “Talking over my problems with friends makes me feel ashamed or foolish.”). Items could be answered on a five-point Likert scale ranging from Never or Almost Never (1) to Always or Almost Always (5). Negative items were reverse coded and a mean score of the 12 items regarding peer attachment was computed such that higher scores indicated higher perceived peer attachment. Cronbach’s alphas ranged between 0.78–0.84 over the four assessments. The IPPA has been shown to have good reliability and validity to assess the quality of the attachment to parents and peers (Gullone and Robinson, 2005; Jewell et al., 2019).

#### 2.3.2. Lottery choice task

The current study adopted a behavioral economic definition of risk, which is defined as the variance between potential outcomes and which allows for the decomposition of risk in a risk and return component (Schonberg et al., 2011). At each yearly assessment, adolescents completed a modified economic lottery choice task (Holt and Laury, 2002) while their blood-oxygen-level dependent (BOLD) responses were recorded. On each trial, participants could choose between two uncertain gambles. Both gambles consisted of a low and a high outcome option, and probabilities for the low and high outcome were the same in both gambles. However, one gamble was riskier because there was a large difference between the high and low outcome (e.g., $3.85 vs $0.10), while the safer gamble had a small difference between the high and low outcome (e.g., $2.00 vs $1.60). Choosing the riskier option would mean a chance of receiving a larger reward at the risk of receiving a much smaller reward, whereas choosing for the safer option would mean that the reward would be similar regardless. The simultaneous presentation of the two gambles was intended to simulate real-world decision options in which individuals are often faced with more than one option, in which some choices are riskier than others.

Participants completed a total of 72 trials in about 25 min. Trials started with a decision phase in which one pair of gambles (a risky and safer option) was presented for four seconds (see Fig. 1). Probabilities and the associated outcomes for these probabilities were presented in purple and orange. Probabilities of each outcome were presented as slices of a pie to increase the numerical understanding of the gamble. The pie consisted of 10 slices with each slice representing 10% probability. Probabilities of the outcomes were the same for each pair of gamble options and were randomly varied across trials (see Supplementary Table S1). Participants could indicate their choice of gamble with a press on a fMRI compatible button box using their left or right hand, which was randomized across participants. After the decision phase a fixation cross was presented with a jitter of 1–3 seconds. Following the fixation period, an outcome screen showed the outcome of the chosen gamble for two seconds. A jittered intertrial interval (ITI) of 1–3 seconds was presented before the start of the next trial. Participants completed six practice trials prior to the fMRI task. Participants received their outcome on four random trials as an additional monetary reward for their participation. Participants were instructed that each trial was independent of the other trial and evenly likely to be selected for payment. As a behavioral indicator of risk taking we computed a risk-preference parameter for each participant using a standard utility power function with higher values indicating higher risk-seeking behavior (see Supplementary Material: Mixed model analyses). Prior studies used the lottery choice task with adolescents and showed that neural responses during this task predicted real-life risky behaviors (Kim-Spoon et al., 2019, 2017).

### 2.4. Control variables

#### 2.4.1. Parental attachment

Prior studies indicated that parental attachment is associated with risk-taking behavior and associated neural responses (Qu et al., 2015). Therefore, to examine the unique effects of peer attachment on neural risk processing, rather than the individual’s general tendency of being attached to people, parental attachment was added to our mixed models (see method section 2.7.1). Parental attachment was assessed using the IPPA (see measures). Adolescents answered the same 12 attachment items for maternal and paternal attachment separately. Negative items were reverse coded and a mean score was computed for maternal and paternal attachment. To create one measure of parental attachment we used the highest score for either maternal or paternal attachment as an index for perceived parental attachment. Cronbach’s alphas ranged between 0.74–0.83 over the four assessments. Similar to other studies, we found that the correlations between peer and parent attachments were low to moderate, ranging between $r = 0.04$ and 0.34 (see
Supplementary Table S2; Armsden and Greenberg, 1987; Raja et al., 1992).

2.4.2. Pubertal development

Prior studies showed that pubertal development is associated with activation in reward-related regions during risk taking (Goddings et al., 2019). Therefore, pubertal development was included to our mixed models (see method section 2.7.1) to investigate neural changes across adolescence over and above the effects that can be explained by pubertal development. Physical pubertal development was assessed using the adolescent self-rating scale of the Pubertal Developmental Scale (Petersen et al., 1988). The questionnaire consisted of five items that assessed physical pubertal developmental such as growth spurts, body hair changes and skin changes. In addition, female adolescents were asked to report on their breast development and menarche, and male adolescents were asked to report on vocal changes and facial hair changes. Items could be answered using a four-point Likert scale, ranging from 1 (no changes) to 4 (changes completed). A mean score of the items was computed as an index for pubertal development. Higher scores indicated more advanced pubertal development. Cronbach’s alphas ranged between 0.44–0.55 over the four assessments. Correlations between pubertal development and age were relatively low, ranging between $r = -0.01$ and 0.22.

2.5. Imaging acquisition

Scans were acquired using a 3.0 T Siemens Tim Trio scanner and a 12-channel head coil (Siemens Healthcare, Erlangen, Germany). Participants could view the screen through a mirror mounted on the head coil. Whole-brain echo-planar imaging (EPI) was used to acquire functional images during the lottery choice task (repetition time (TR) = 2000 ms, echo time (TE) = 30 ms, field of view (FOV) = $220 \times 220$ mm, 64 x 64 grid, voxel size = $3.4 \times 3.4 \times 3.4$ mm, flip angle = 90°, 34 axial slices, and 4.0 mm slice thickness). Hyperangulated slices were acquired at 30 degrees from anterior-posterior commissure. For within-subject registration, a whole-brain T1-weighted anatomical scan was acquired using high-resolution magnetization prepared rapid acquisition gradient echo sequence (192 slices, TR = 1200 ms, TE = 2.66 ms, FOV = $245 \times 245$ mm, voxel size = $1 \times 1 \times 1$ mm). There were 141 valid scans obtained at T1, 136 at T2, 126 at T3, and 129 at T4 (see Supplementary Table S3). Reasons for excluding scans include not meeting MRI safety criteria, excessive head motion (>3 mm), technical errors, or imaging artifacts.

2.6. fMRI preprocessing and analysis

Functional images were preprocessed and analyzed using Statistical Parametric Mapping version 8 (SPM8: Wellcome Centre for Human Neuroimaging, London, UK). Functional images were realigned using six rigid body transformation parameters. To correct for the timing differences in slice acquisition, slice time correction was performed. The mean functional image was co-registered with the anatomical image. The T1-weighted anatomical image was segmented and co-registered to the Montreal Neurological Institute template using the unified segmentation approach with a resampling rate of $3 \times 3 \times 3$ mm voxel size (Ashburner and Friston, 2005). These co-registration parameters were used to normalize the functional images. Finally, normalized images were smoothed with a Full Width Half Maximum Gaussian kernel of 6 mm.

For each yearly assessment, a first-level general linear model (GLM) was conducted which included one regressor for the decision period and one regressor for the outcome period with a fixed duration of four and two seconds, respectively. A parametric regressor was included that coded for the magnitude of risk of a chosen gamble on each trial to examine which brain regions were sensitive to the magnitude of risk. To this end, we quantified the amount of risk associated with each gamble during the lottery choice task by computing the coefficient of variation (CV; see Supplementary Material). The CV is a scale-free metric and has been shown to be superior in explaining choice behavior compared to other economic measures of risk (i.e., standard deviation or variance) because outcomes are coded by the relative risk as opposed to the absolute outcome (Bach et al., 2017; Weber et al., 2004). The CV is computed by dividing the standard deviation of the chosen gamble by the expected value (i.e., probability-weighted outcome) of that gamble, with higher levels of CV corresponding to a higher variance between the high and low outcome on a gamble and thus higher risk (see Supplementary Table S1). The CV of the chosen gamble on each trial was added as parametric regressor to the GLM. Another parametric regressor was presented with the same probabilities of winning a high or low outcome. However, gambles differed on the variance in low and high outcome. That is, on riskier gambles (right gamble) the difference between the high and low outcome was larger ($3.85 - 0.10$) than on the safer gambles (left gamble, $2.00 - 1.60$). Outcomes and probabilities were presented in corresponding colors (purple and orange). The decision phase was followed by a fixation screen after which the outcome on the chosen gamble was shown. Trials ended with a jittered intertrial interval (ITI).
added to the GLM that coded for high and low outcomes during the outcome phase to assess outcome processing. Furthermore, one regressor for the button press and six motion regressors were included. These regressors of interest and no interest were included to appropriately characterize the error term in the neural model. Regressors were convolved with a canonical hemodynamic response function. Data were prewhitened to remove temporal autocorrelation (AR1) and high-pass filtered at 128 Hz.

First-level images from each year assessing neural responses associated with the CV of the chosen gamble during the decision were submitted to whole-brain random-effect analysis separately. On each assessment, the CV of the chosen gamble was significantly associated with the BOLD responses in the insula and dACC, such that choosing riskier gambles was related to higher BOLD responses in the insula and dACC (Fig. 2). All statistical inferences were made at a cluster-corrected threshold of \( p < .001 \) with a Family-Wise Error (FWE) correction, with an initial uncorrected threshold of \( p < .001 \). Results indicated that the bilateral insula and dACC belonged to the largest regions, with greatest \( t \)-values, among the regions activated in response to the differing levels of risk in our task (see Supplementary Tables S4-7). Additionally, they were the most consistently activated brain regions across all four waves in our longitudinal assessments. In contrast, activation in the ventral striatum, orbitofrontal cortex, and lateral prefrontal cortex was not consistent across assessments. This is not surprising given the large body of literature in both adolescents (e.g., van Duijvenvoorde et al., 2015) and adults (e.g., Mohr et al., 2010) indicating that the insula and dACC are key regions consistently reported to represent economic risk in the brain. Accordingly, only the bilateral insula and dACC were considered in the main analyses. For consistent comparison of each yearly assessment, eigenvariate values from the bilateral insula and the dACC were extracted using a standard neuroanatomical atlas for functional neuro-imaging data (Automated Anatomical Labeling (AAL) (Tzourio-Mazoyer et al., 2002) using SPM (see Supplementary Figure S1). Once the anatomical ROIs were extracted, we performed our statistical analyses.

### 2.7. Statistical analyses

#### 2.7.1. Mixed model building procedure

We used a mixed model approach in R (R core Team, 2018) using the nlme package (Pinheiro et al., 2018). Mixed models are appropriate for nested data such as longitudinal data as they allow for the estimation of average starting points and trajectories in the sample (i.e., fixed intercepts and slopes) while at the same time allowing for individual variation in starting points and trajectories over time (i.e., random intercepts and slopes). Data was missing completely at random (Little MCAR test: \( \chi^2 = 586.11, \text{df} = 666, p = .981 \)) and therefore multiple imputation was used with the MICE package in R (van Buuren and Groothuis-Oudshoorn, 2011). Missing data on neural risk processing, peer attachment and covariates were imputed. Imputing data that are missing completely at random has been shown to not bias parameter estimates (Ambler et al., 2007; Donders et al., 2006; Little, 1992; Rubright et al., 2014). Moreover, multiple imputations have been shown to outperform deletion methods (Enders et al., 2016; Newman, 2003).

Our first goal was to assess the development of neural risk processing in the insula and dACC across adolescence using a formal fitting approach (Braams and Crone, 2017; Braams et al., 2015). Neural risk processing in the insula and dACC were fitted separately using a maximum likelihood estimation (for the full model see supplementary material). First, a null model that included both a fixed and random intercept was fitted. Next, we compared the null model against a set of models assessing developmental changes (i.e., linear, quadratic, cubic) by adding age (mean-centered) as a polynomial predictor. Models were compared using Akaike Information Criterion (AIC; Akaike, 1974) and the Bayesian Information Criterion (BIC; Schwarz, 1978) values. Lower AIC and BIC values indicate a better model fit. Following identification of the best fitting developmental model, a random slope for age was added to examine whether there were significant individual differences in development of neural risk processing. All models showed significant improvement with the addition of a random slope \( (p < .001) \). We also examined sex differences in the development of neural risk processing by adding sex as a main effect and interaction effect to the best fitting polynomial age model. Each adolescent’s sex was dummy-coded with males as 0 and females as 1. Finally, given the influence that puberty has on brain development (Blakemore et al., 2010), follow-up analyses included time-varying levels of pubertal development in the best fitting developmental model to evaluate the independent contribution of age on neural risk processing development.

Our second goal was to examine how peer attachment across ages 13–17 years was associated with the development of neural risk processing in the insula and dACC. We first fitted a model that included sex and the best fitting age model (e.g., linear, quadratic or cubic) of insula or dACC risk processing. We then tested for main effects of peer attachment. Next, we fitted a model assessing associations over time (peer attachment X age), a model assessing sex differences (peer attachment X sex) and a model that included both two-way interactions. Finally, to rule out higher-order associations we also fitted a model with
the three-way interaction among peer attachment, age, and sex (peer attachment X age X sex). Model fit was assessed by evaluation of AIC. BIC values are reported for completeness but are not used in model selection as BIC values may perform worse with complex models such as models including interactions (Vrieze, 2012). Follow-up analyses included time-varying levels of parental attachment and pubertal development, to evaluate the independent contribution of peer attachment on neural risk processing and above perceived parental attachment and pubertal development. Given that our models assessing sex-specific effects were exploratory, results of these analyses are reported at a significance level of \( p < .05 \) (Bender and Lange, 2001).

### 2.7.2. Intra-class correlation coefficient (ICC)

Mixed model analyses require that observations of longitudinal data are sufficiently nested in individuals and can be assessed with an intra-class correlation coefficient (ICC). ICC values were assessed using two-way mixed effects models with absolute agreement in SPSS, version 25. All ICC values of our study variables are above 0.10 indicating satisfactory nesting of data within individuals (see Table 1) (Gicchetti and Sparrow, 1981).

### 3. Results

#### 3.1. Descriptive statistics

Descriptive statistics are shown in Table 1. Females scored higher on perceived peer attachment on T1 but did not differ from males on T2-T4. Results of the whole-brain analysis are presented in Supplementary Tables S4 - S7 and correlations between study variables are shown in Supplementary Table S2. For descriptive purposes, developmental changes of peer attachment, behavioral performances on the lottery choice task (i.e., risk-preference) and the covariates of pubertal status and attachment were assessed using mixed model analyses (see Supplementary Tables S8-9). The development of peer attachment differed between males and females (\( B_{sex \times age} = -0.09, SE = 0.03, p = .002 \)). The level of peer attachment was stable for females (\( B_{age\text{linear}} = -0.02, SE = 0.02, p = .511 \)), whereas males started at a lower level (see Table 1) but increased (\( B_{age \text{linear}} = 0.07, SE = 0.02, p < .001 \)) in their perceived peer attachment to levels similar to females by age 17 years. No significant sex differences were found in the patterns of development of behavioral risk-preference. Rather, risk-preference on the lottery choice task showed a linear decrease with age for both males and females (\( B_{age \text{linear}} = -0.12, SE = 0.03, p < .001 \)). Pubertal development increased with age and stabilized at age 17 (\( B_{age \text{linear}} = 0.19, SE = 0.01, p < .001 \); \( B_{age \text{quadratic}} = -0.03, SE = 0.01, p < .001 \)) with higher levels of pubertal development for females than males across ages 13–17 years (\( B_{sex} = 0.64, SE = 0.04, p < .001 \)). Finally, the level of parental attachment was stable over time (\( B_{age \text{linear}} = -0.01, SE = 0.02, p = .644 \)) and did not differ between males and females.

#### 3.2. Development of neural risk processing

AIC and BIC values for the mixed models examining the development of insula and dACC risk processing are listed in Table 2. Results showed that the development of insula and dACC risk processing across ages 13–17 years was best described by a linear growth pattern. As shown in Table 3, insula and dACC risk processing increased linearly across the measured period (right insula: \( B_{age} = 0.18, SE = 0.03, p < .001; \) left insula: \( B_{age} = 0.15, SE = 0.02, p < .001; \) and dACC: \( B_{age} = 0.18, SE = 0.03, p < .001 \)). No main effect or interaction with sex was found in the development of insula and dACC risk processing. Controlling for time-varying levels of pubertal development did not change these results (see Supplementary Table S10).

#### 3.3. Associations of peer attachment with neural risk processing

Linear mixed models were fitted to examine the association between

### Table 1

Means (M) and Standard Deviations (SD) of Peer Attachment, Behavioral and Neural Correlates of Risk Processing, and Control Variables.

|       | T1 Male |       | SE | T1 Female |       | SE | T2 Male |       | SE | T2 Female |       | SE | T3 Male |       | SE | T3 Female |       | SE | T4 Male |       | SE | T4 Female |       | SE | ICC (95% CI) |
|-------|---------|-------|----|----------|-------|----|---------|-------|----|----------|-------|----|---------|-------|----|----------|-------|----|---------|-------|----|----------|-------|----|----------|-------|----|
| Age   | 14.19   | 0.52  | 14.06| 0.55     | 15.13 | 0.52| 14.98  | 0.57     | 16.15 | 0.54 | 16.01 | 0.57 | 17.08 | 0.53 | 16.91 | 0.56 | .73 | (.65–.79) |
| Peer attachment | 3.88 | 0.51 | 4.12 | .59 | 3.99 | 0.45 | 4.14 | 0.58 | 4.08 | 0.45 | 4.14 | 0.50 | 4.11 | 0.41 | 4.07 | 0.67 | .46 | (.31–.58) |
| Risk-preference | 0.72 | 0.49 | 0.88 | 0.93 | 0.54 | 0.48 | 0.66 | 0.87 | 0.59 | 0.91 | 0.42 | 0.39 | 0.37 | 0.33 | 0.45 | 0.90 | .46 | (.31–.58) |
| Right insula | 0.03 | 0.05 | 0.03 | 0.05 | 0.46 | 0.67 | 0.47 | 0.58 | 0.75 | 1.04 | 0.54 | 0.74 | 0.61 | 1.09 | 0.61 | 0.77 | .46 | (.32–.58) |
| Left insula | 0.03 | 0.05 | 0.04 | 0.05 | 0.41 | 0.68 | 0.38 | 0.52 | 0.60 | 0.99 | 0.42 | 0.78 | 0.53 | 1.15 | 0.56 | 0.74 | .49 | (.35–.61) |
| dACC   | 0.04   | 0.05  | 0.03 | 0.04 | 0.44 | 0.75 | 0.39 | 0.54 | 0.73 | 1.05 | 0.57 | 0.87 | 0.55 | 1.04 | 0.68 | 0.94 | .81 | (.64–.88) |
| Control variables | 2.62 | 0.51 | 3.24 | 0.44 | 2.83 | 0.38 | 3.45 | 0.44 | 2.99 | 0.36 | 3.67 | 0.31 | 3.19 | 0.40 | 3.82 | 0.33 | .86 | (.82–.89) |
| Parental attachment | 4.10 | 0.48 | 4.25 | 0.48 | 4.02 | 0.56 | 4.16 | 0.58 | 4.13 | 0.50 | 4.10 | 0.61 | 4.08 | 0.53 | 4.14 | 0.59 | .64 | (.64–.88) |

Note: Bold entries indicate significant differences between males and females at \( p < .05 \). T1-4 = time point 1–4.

ICC = intra-class correlation coefficient; CI = confidence interval; dACC = dorsal anterior cingulate cortex; PD = pubertal development.
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Across Ages 13 and sex, suggesting that the strength of two-way interaction between attachment X sex X age). This model did not improve the model fit. Moreover, allowing the interaction between peer attachment and risk processing to vary over time (peer attachment X age) did not improve model fit. However, including the interaction between peer attachment and sex improved model fit. To examine sex-specific changes over time in this association, we also assessed a model including the three-way interaction between peer attachment, sex, and age (peer attachment X sex X age). This model did not improve the model fit relative to the model including the interaction between peer attachment and sex, suggesting that the strength of two-way interaction between peer attachment and sex did not significantly change from 13 to 17 years. Therefore, for the right insula, the model including the two-way interaction between peer attachment and sex was selected as the best-fitting model. For the left insula and dACC, including peer attachment or higher order interactions did not improve model fit. Final model parameters of the models including the interaction between peer attachment and sex are shown in Table 5.

For right insula risk processing, we found a significant interaction between peer attachment and sex ($B_{attachment*sex} = -0.21$, $SE = 0.10$, $p = .034$; see Fig. 3). We performed simple slope analysis of this significant interaction to probe the direction of effect under the moderator variable (i.e., sex) (Holmbeck, 2002). In other words, we examined the magnitude and direction of the association between right insula risk processing and peer attachment separately for females and males. This simple slope analysis showed a significant positive association between peer attachment and neural risk processing only for males ($B_{attachment} = 0.19$, $SE = 0.09$, $p = .039$); there was no significant association between peer attachment and right insula processing for females ($B_{attachment} = -0.03$, $SE = 0.06$, $p = .659$). Our results remained significant after controlling for time-varying levels of pubertal development and parental attachment (see Supplementary Table S11). As can be seen in Table 5, there was no significant main effect or interaction for peer attachment with the left insula and dACC.

Results thus showed that for male adolescents the level of peer attachment was associated with neural risk processing (particularly in right insula) and the strength of this association was stable from ages 13–17 years. As supplemental analysis, we also assessed associations between peer attachment and behavioral performances on the task (i.e., risk-preference). No significant main effect of peer attachment or interaction effect with sex was found (see Supplementary Table S12).

4. Discussion

The aims of this longitudinal study were to examine the development of neural risk processing across adolescence, and to examine its associations with peer attachment. Results showed that the insula and dACC were consistently activated during neural risk processing across adolescence and showed a linear increase from ages 13 to 17 years. Furthermore, there was a positive association between peer attachment and right insula risk processing for male adolescents, and the strength of this association was stable across ages 13–17 years. The association between peer attachment and neural risk processing was not significant for female adolescents.

The current results present the first insight into the development of insula and dACC risk processing as well as providing evidence of longitudinal associations between peer attachment and right insula risk processing, particularly for male adolescents. With respect to our finding on the development, specifically linear increases in insula and dACC risk processing across ages 13–17 years, this finding is in line with previous cross-sectional studies that compared different age groups and demonstrated higher activation during adolescence compared to childhood (Paulsen et al., 2012; van Duijvenvoorde et al., 2015). Specifically, Paulsen et al. (2012) reported linear age-related increases in insula and dACC activation during risk processing when comparing children, adolescents, and adults; whereas van Duijvenvoorde et al. (2015) reported heightened insula and dorsal medial prefrontal cortex activation in late adolescents compared to children and adults. Our finding clarifies the developmental pattern of insula and dACC activation during risk processing by examining within-person changes across adolescence. The found increase also concurs with the theory that the prefrontal cortex matures gradually during adolescence and gains control over adolescent risk-taking decision making (Casey et al., 2008; Steinberg, 2007). However, we note that some previous cross-sectional studies showed lower dACC activation during risk taking in adults compared to adolescents and children (van Leijenhorst et al., 2006, 2010). These studies have either used tasks involving one risky gamble per trial or tasks in which the participant was asked to decide between a certain amount or a risky gamble (e.g., van Leijenhorst et al., 2006, 2010). Whereas, in the current study, both high- and low-risk gambles were presented simultaneously to simulate decision options in the real world. That is, in real world risky situations, individuals are often faced with two options, in which one option is relatively riskier than the other. Accordingly, both options have potential consequences that are often judged relative to one another. This paired presentation may have potentially increased the salience of high-risk choices.

Table 3

| Parameter Estimates (Regression Coefficient \(b\), Standard Error \(SE\) and Significance Level \(p\)) for the Best Fitting Developmental Model of Neural Risk Processing Across Ages 13 – 17 Years. |
|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
|                | Right insula    | Left insula     | dACC            |                |                |                |                |                |                |
|                | \(b\)           | \(SE\)          | \(p\)           | \(b\)          | \(SE\)          | \(p\)           | \(b\)          | \(SE\)          | \(p\)           |
| Intercept      | 0.44            | 0.04            | <.001           | 0.38           | 0.04            | <.001           | 0.45           | 0.04            | <.001           |
| Age (linear)   | 0.18            | 0.03            | <.001           | 0.15           | 0.03            | <.001           | 0.18           | 0.02            | <.001           |

Note. All models include a random slope for age. dACC = dorsal anterior cingulate cortex.

Table 4

AIC and BIC values of the Models Assessing Associations of Peer Attachment with Neural Risk Processing and Risk-Preference.

<table>
<thead>
<tr>
<th>Model</th>
<th>Age + Sex</th>
<th>Age + Sex + peer attachment</th>
<th>Peer attachment X Age</th>
<th>Peer attachment X Sex</th>
<th>Peer attachment X Age + Peer attachment X Sex</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>AIC</td>
<td>BIC</td>
<td>AIC</td>
<td>BIC</td>
<td>AIC</td>
</tr>
<tr>
<td>Right insula</td>
<td>1360</td>
<td>1386</td>
<td>1361</td>
<td>1399</td>
<td>1363</td>
</tr>
<tr>
<td>Left insula</td>
<td>1388</td>
<td>1415</td>
<td>1390</td>
<td>1421</td>
<td>1392</td>
</tr>
<tr>
<td>dACC</td>
<td>1441</td>
<td>1468</td>
<td>1443</td>
<td>1474</td>
<td>1444</td>
</tr>
<tr>
<td>Risk-preference</td>
<td>1423</td>
<td>1459</td>
<td>1425</td>
<td>1466</td>
<td>1427</td>
</tr>
</tbody>
</table>

Note. Preferred models are marked in bold. AIC = Akaike Information Criterion, BIC = Bayesian Information Criterion.
Besides the neural development of risk processing, our finding of an association between higher peer attachment and greater neural risk processing in the right insula across ages 13–17 years, albeit only among males, is also noteworthy. Previous work showed that high peer attachment buffered the effect of peer conflict on the socio-emotional system (e.g., striatum and medial prefrontal cortex) during risk taking (Telzer et al., 2015), suggesting the protective effect of peer attachment against the negative effect of peer conflict on reward sensitivity. Prior neuroimaging studies focusing on risk sensitivity have demonstrated a crucial involvement of the right insula in decision risk—i.e., risk processing during or before choice (Mohr et al., 2010; van Duijvenvoorde et al., 2015) and risk-averse behavior (Van Duijvenvoorde et al., 2015). Our finding of higher right insula activation for adolescents high on peer attachment possibly indicates that the insula was effectively working to lead decision makers to exhibit caution and thoroughly evaluate risky options. The current study extends prior work by providing evidence that higher peer attachment is associated with increased neural processing of risk information. The findings seem to suggest the beneficial nature of peer attachment on brain function—i.e., prompting adolescents away from (rather than towards) risky choices.

Importantly though, the significant association between peer attachment and right insula processing was only found among male adolescents. Most previous studies that examined peer influences on adolescent brain functioning were based on relatively small samples and did not test sex differences in these associations. Our results on male specific activation is not in line with the study by Braams et al. (2014) who explored sex differences. That study found a positive association between friendship quality and ventral striatum activation when winning a gamble for a friend for females, but not males (Braams et al., 2014). Differences between the current study and the prior study may be explained by differences in measured neural processes—e.g., risk processing versus reward processing (d’Acremont and Bossaerts, 2008; van Duijvenvoorde et al., 2015) as well as existing sex differences in the nature of attachment to peers. That is, female adolescents value peer intimacy more and seek peer support more, whereas male adolescents prefer spending time with their peers based on common interests (Gorrese and Ruggieri, 2012; Rose and Rudolph, 2006). Thus, winning money for a friend on a gamble may be perceived as more rewarding for females (Braams et al., 2014) due to the value they place on their peer network as secure base. In contrast, risk-taking behavior may be a way to gain recognition from peers and be part of a peer group particularly for male adolescents (Michael and Ben-Zur, 2007; Piko, 2000). Indeed, a prior study showed that social exclusion by peers increased levels of risk taking particularly for adolescents who were more susceptible to peer influences, and this was negatively related to activation in the lateral prefrontal cortex, a region important for exerting behavioral control (Eshel et al., 2007; Peake et al., 2013). Thus, male adolescents who felt greater psychological security provided by peers exhibited more careful brain processing when evaluating risky options, whereas adolescents lacking good quality relationships may be more likely to take risks as a part of their efforts to overcome negative feelings resulting from their low acceptance by peers (Brady et al., 2009).

Our data revealed that the effect of peer attachment was only observed for risk-related neural processing, but not for the behavioral measure of risky decision-making (i.e., risk preference). The reason for the association of peer attachment with neural and not behavioral measures may be due to the fact that peer attachment is related to the processing of risk, which is captured by the neural activation, rather than the outcome, which is reflected through behavioral choices. That is, adolescents evaluated risk information related to options before making a choice. Moreover, laboratory tasks may be limited in capturing real-world behavioral responses, but risk-related neural processing may be able to more accurately represent individual differences in neurobiological processes (Richards et al., 2013). In addition, although we were not able to test this, we note that associations between neural risk processing and peer attachment may differ depending on the nature of the

Table 5

<table>
<thead>
<tr>
<th></th>
<th>Right insula</th>
<th></th>
<th>Left insula</th>
<th></th>
<th>dACC</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>b</td>
<td>SE</td>
<td>p</td>
<td>b</td>
<td>SE</td>
</tr>
<tr>
<td>Intercept</td>
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<td>0.05</td>
<td>&lt;.001</td>
<td>0.41</td>
<td>0.05</td>
</tr>
<tr>
<td>Age (linear)</td>
<td>0.17</td>
<td>0.03</td>
<td>&lt;.001</td>
<td>0.15</td>
<td>0.03</td>
</tr>
<tr>
<td>Sex</td>
<td>–0.06</td>
<td>0.08</td>
<td>.376</td>
<td>–0.05</td>
<td>0.07</td>
</tr>
<tr>
<td>Peer attachment</td>
<td>0.19</td>
<td>0.06</td>
<td>.015</td>
<td>0.09</td>
<td>0.08</td>
</tr>
<tr>
<td>Peer attachment X Sex</td>
<td>–0.22</td>
<td>0.10</td>
<td>.034</td>
<td>–0.13</td>
<td>0.11</td>
</tr>
</tbody>
</table>

Note. All models include a random slope for age. Male is coded as 0, female as 1. dACC = dorsal anterior cingulate cortex.

Fig. 3. Associations between peer attachment and neural risk processing for males and females in the right insula. Shades around the regression lines represent the 95% confidence interval. *p < .05.
peer interactions. That is, adolescents tend to select and associate with similarly behaving peers, and peer socialization through deviant peer association is known to be one of the strongest predictors of adolescent delinquent behaviors (Dishion and Owen, 2002; Dishion and Patterson, 2016). Indeed, prior studies showed that interacting with a risk-promoting peer was related to higher risk-taking behaviors, whereas interacting with a cautious peer was associated with lower risk-taking behaviors (see Dishion and Patterson, 2016 for review). Moreover, a recent study reported a longitudinal association between adolescents’ affiliation with substance using peers and lowered insula activation during risk processing (Kim-Spoon et al., 2019). As such, future studies are needed to investigate how the nature of peer attachment (e.g., deviant or low-risk peers) may relate to adolescents’ neural risk processing.

An important strength of this study is the inclusion of adolescents from an Appalachian region of southwestern Virginia. This region is a distinct geographic and cultural area of the United States and adolescents from these regions are often from low household incomes (Census, 2017). Moreover, adolescents from these rural regions show relatively higher incidences of health risk behaviors such as substance use or risky sexual behavior (Moreland et al., 2013; Wewers et al., 2006), providing implications for preventive intervention efforts. Another methodological strength is that peer attachment was associated with neural risk processing even when controlling for parental attachment. This demonstrates that the level of attachment to peers uniquely contributes to neural risk processing across adolescence rather than being attached to people (e.g., peers and parents) in general.

Despite its strengths, some limitations of this study must be noted. First, more research is necessary to determine the generalizability of our findings and possible variations to these findings due to different cultural contexts. Second, in the current study risk processing was defined as the variance in potential outcomes, and the lottery choice task did not include gambles where potential losses were possible. Although the absence of a high rewarding gain could be interpreted as a loss in our study, the current results may not be generalizable to all forms of risk taking, as some forms of risk-taking behavior in everyday life include potential harmful outcomes (Schonberg et al., 2011). Third, the associations between peer attachment and neural risk processing were correlational, so causal effects could not be inferred. Lastly, our analyses on sex differences were exploratory. Future research utilizing even larger samples is necessary to replicate our findings and gain a better understanding of how peer attachment may relate to neural risk processing for both male and female adolescents.

To our knowledge, this is the first prospective longitudinal study demonstrating how peer attachment is related to the development of brain functioning related to risk processing throughout adolescence. We observed that neural sensitivity to risk in the insula and dACC increased from 13 to 17 years, and that individual differences in right insula risk processing associated with levels of peer attachment, but only for males. The found increases in neural sensitivity for risk suggest that adolescents become more prudent decision makers across adolescence, according to the brain maturation that facilitates the consideration of the potential riskiness in their decisions. In addition, the found association between peer attachment and neural risk processing for males has important implications for preventative interventions as adolescent males have been shown to engage in higher levels of risk behaviors compared to female adolescents (Moffitt and Caspi, 2001). Also, lower neural sensitivity to risk has been shown to predict later health-risk behaviors such as substance abuse (Kim-Spoon et al., 2019). Present findings demonstrate the potential role of peer relationships in adolescent risk taking, suggesting that greater risk sensitivity in the male adolescent brain may be a mechanism by which supportive peer relationships could demote risky behaviors during adolescence.

Disclaimer

The views and opinions expressed in this article are those of the authors and do not necessarily reflect the official policy or position of any agency of the U.S. government.

Declaration of Competing Interest

The authors declare no conflict of interest.

Acknowledgments

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Appendix A. Supplementary data

Supplementary material related to this article can be found, in the online version, at doi:https://doi.org/10.1016/j.dcn.2020.100772.

References


